

Quenched localisation in the Bouchaud trap model with regularly varying traps

By

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Abstract

This article describes the quenched localisation behaviour of the Bouchaud trap model on the integers with regularly varying traps. In particular, it establishes that for almost every trapping landscape there exist arbitrarily large times at which the system is highly localised on one site, and also arbitrarily large times at which the system is completely delocalised.

§ 1. Introduction

The Bouchaud trap model (BTM) was introduced in [3] as a simple way of investigating the evolution of a physical system – particularly a spin glass – through a sequence of meta-stable states. A distinctive feature of the systems considered by Bouchaud is that they exhibit the phenomena of *ageing*, meaning that we can tell how long the system has been running by observing its present state. At least in one-dimensional versions of the model where there is a limited number of accessible sites, this property of ageing turns out to be intrinsically related to *localisation*, namely that at certain times we can predict with high probability the state of the system [2]. It is the goal of this article to study this latter property in the special case of the BTM on the integers with regularly varying traps.

We now introduce the model of study, following the notation of [5]. First, define a trapping landscape $\sigma = (\sigma_x)_{x \in \mathbb{Z}}$, which is a collection of independent and identically

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distributed (i.i.d.) strictly-positive random variables, built on a probability space with probability measure \mathbf{P} . Conditional on σ , the dynamics of the BTM are given by a continuous-time \mathbb{Z} -valued Markov chain $X = (X_t)_{t \geq 0}$, started from the origin, with transition rates

$$(1.1) \quad w_{x \rightarrow y} = \begin{cases} \frac{1}{2\sigma_x}, & \text{if } y \sim x, \\ 0, & \text{otherwise,} \end{cases}$$

where $y \sim x$ means that x and y are nearest neighbours in \mathbb{Z} . We denote the law of X conditional on σ , the so-called ‘quenched’ law of the BTM, by P_σ . Throughout the article we will suppose that the trap distribution σ_0 satisfies

$$(1.2) \quad \mathbf{P}(\sigma_0 \geq u) = u^{-\alpha}, \quad \forall u \geq 1,$$

for some constant $\alpha \in (0, 1]$. Whilst this is a strict assumption, we believe that, after making suitable adaptations to the arguments, one could deduce the same results under certain weaker assumptions. For instance, in the case $\alpha \in (0, 1)$, it should be sufficient that the tail of the distribution of σ_0 be regularly varying with the same index. We exclude the parameter range $\alpha > 1$ since in this regime the BTM does not exhibit localisation, as explained below.

Specifically, the aim of this article is to establish the following quenched localisation behaviour of the BTM on the integers with regularly varying traps. On the one hand, we prove that for almost every trapping landscape there are arbitrarily large times at which the system is highly localised on one site, a site that can be described explicitly in terms of the trapping landscape. Conversely, we show that for almost every trapping landscape there are also arbitrarily large times at which the BTM is completely delocalised, i.e. no single site carries a prescribed amount of probability.

Theorem 1.1. *For the BTM on the integers with a trapping distribution satisfying (1.2) for some $\alpha \in (0, 1]$, it \mathbf{P} -a.s. holds that*

$$(1.3) \quad \liminf_{t \rightarrow \infty} \sup_{x \in \mathbb{Z}} P_\sigma(X_t = x) = 0, \quad \limsup_{t \rightarrow \infty} \sup_{x \in \mathbb{Z}} P_\sigma(X_t = x) = 1.$$

We now compare this result with previous studies of the quenched behaviour of the BTM that have appeared in the literature. In [7] (see also [8]) it was established that, for $\alpha \in (0, 1)$, the probability mass function of the BTM exhibits quenched localisation, in the sense that

$$\sup_{x \in \mathbb{Z}} P_\sigma(X_t = x) \not\rightarrow 0$$

as $t \rightarrow \infty$. Theorem 1.1 strengthens the above localisation result by demonstrating that the supremum (indeed the ℓ_p -norm, for any $p > 1$) of the probability mass function of the BTM fluctuates infinitely often between the bounds of 0 and 1.

One interesting consequence of Theorem 1.1 is to demonstrate a relatively sharp transition in the quenched behaviour of the BTM between the homogenised regime and a regime of strong disorder, as the tail of the trap distribution gets heavier. Recall that if σ_0 has finite mean (and in particular if (1.2) holds for some $\alpha > 1$), the BTM homogenises \mathbf{P} -a.s. In other words the BTM, rescaled diffusively, converges in distribution to Brownian motion \mathbf{P} -a.s. In Section 4 we check that this, in turn, implies that, \mathbf{P} -a.s., as $t \rightarrow \infty$,

$$(1.4) \quad \sup_{x \in \mathbb{Z}} P_\sigma (X_t = x) \rightarrow 0.$$

By contrast, Theorem 1.1 demonstrates that, if $\alpha \leq 1$, there are arbitrarily large times, \mathbf{P} -a.s., at which the probability mass function of the BTM is in a maximally disordered state. Whether (1.3) remains true for *any* σ_0 with infinite mean (perhaps under suitable regularity conditions) is an interesting question which we leave open.

We also note that the recent work [5] investigated the result corresponding to Theorem 1.1 in the case of σ_0 with a slowly varying tail at infinity (roughly this is the $\alpha = 0$ case; see [6, Theorem 1.9]). In particular, it was demonstrated that in the one-sided case (i.e. restricting the BTM to the positive integers), there exist distributions of σ_0 such that

$$\liminf_{t \rightarrow \infty} \sup_{x \in \mathbb{Z}^+} P_\sigma (X_t = x) = \frac{1}{N}, \quad \limsup_{t \rightarrow \infty} \sup_{x \in \mathbb{Z}^+} P_\sigma (X_t = x) = 1$$

for each $N \in \{2, 3, \dots\}$. We suspect that the equivalent result also holds true in the two-sided case, for N restricted to $\{3, 4, \dots\}$. Whether there exist trap distributions σ_0 for which

$$\liminf_{t \rightarrow \infty} \sup_{x \in \mathbb{Z}} P_\sigma (X_t = x) = p$$

for arbitrary $p \in [0, 1/3]$ (or $p \in [0, 1/2]$ in the one-sided case) is also an interesting open question.

Our approach to establishing the limsup part of Theorem 1.1, which we do in Section 2, largely follows the argument in [5]. On the other hand, our argument for the liminf part in Section 3 makes use of heat-kernel estimates, which is quite different from the approach taken for the equivalent bounds in [5].

Finally, we remark that as a by-product of our argument we establish bounds on the almost sure fluctuations in the sum/max ratio of i.i.d. sequences of random variables with common distribution σ_0 . Again, one might compare to the integrable case, in which the liminf will diverge, and the slowly varying case, in which the limsup can be finite.

Theorem 1.2. *Assume σ_0 satisfies (1.2) for some $\alpha \in (0, 1]$. Denote by S_n and*

M_n the sum and maximum respectively of the partial sequence $(\sigma_i)_{1 \leq i \leq n}$

$$S_n := \sum_{1 \leq i \leq n} \sigma_i, \quad M_n := \max_{1 \leq i \leq n} \sigma_i.$$

Then it \mathbf{P} -a.s. holds that

$$\liminf_{n \rightarrow \infty} \frac{S_n}{M_n} = 1, \quad \limsup_{n \rightarrow \infty} \frac{S_n}{M_n} = \infty.$$

§ 2. Localisation on a single point

The aim of this section is to prove that at arbitrarily large times the BTM is highly localised, that is, to prove the lim sup part of Theorem 1.1. Our approach is to show that certain favourable configurations of the trapping landscape occur infinitely often \mathbf{P} -a.s., and moreover, when such favourable configurations arise, the BTM is highly localised on a single site at a certain time.

We first introduce an $\varepsilon \in (0, 1)$ that will act as our error threshold. Unless explicitly stated, ε will remain fixed throughout this section, and as such we shall suppress the explicit dependence of other notation on ε . To define the favourable configurations, we introduce the scales

$$(2.1) \quad a_n := \lfloor e^{2n \log n} \rfloor, \quad b_n := \varepsilon^{-1} \lfloor e^{2n \log n} \rfloor.$$

Note that we have chosen a_n specifically so that the ratio $a_{n-1}/a_n \sim n^{-2} \rightarrow \infty$ is a summable sequence. Further, recalling the notation for the one-sided sum and maximum processes S_n and M_n from the statement of Theorem 1.2, we introduce the two-sided sum process

$$\bar{S}_n := \sum_{-n \leq i \leq n} \sigma_i,$$

and define the events, for $n \in \mathbb{N}$,

$$\mathcal{E}_n := \left\{ M_{a_n} > \varepsilon^{-3} a_n^{1/\alpha} \ell_\alpha(a_n), \bar{S}_{b_n} - M_{a_n} < 3\varepsilon^{-1} a_n^{1/\alpha} \ell_\alpha(a_n) \right\},$$

where

$$\ell_\alpha(n) := \begin{cases} 1, & \text{if } \alpha \in (0, 1), \\ \log n, & \text{if } \alpha = 1, \end{cases}$$

is a logarithmic correction in the case $\alpha = 1$.

We show that, for each $\varepsilon \in (0, 1)$, the events \mathcal{E}_n occur infinitely often \mathbf{P} -a.s. (see Lemma 2.2). Moreover, we show that, when the event \mathcal{E}_n occurs, at the time

$$t_n := \varepsilon^{-2} a_n^{1+1/\alpha} \ell_\alpha(a_n)$$

the BTM is completely localised, up to an ε -dependent error, on the site achieving the maximum M_{a_n} (see Corollary 2.4), a site that we shall denote by

$$x_n := \arg \max_{0 \leq i \leq a_n} \sigma_i.$$

To establish that the events \mathcal{E}_n occur infinitely often, we start by proving the corresponding result for a closely related sequence of independent events. In particular, define the collections of intervals $(I_n)_{n \geq 0}$ and $(J_n)_{n \geq 0}$, by setting $I_0 := (0, a_0]$, $J_0 := [-b_0, 0] \cup (a_0, b_0]$, and, for $n \geq 1$,

$$I_n := (b_{n-1}, a_n], \quad J_n := [-b_n, -b_{n-1}) \cup (a_n, b_n].$$

Note that $\cup_{n=0}^{\infty} (I_n \cup J_n) = \mathbb{Z}$, and also I_n, J_n , $n \geq 1$, are all disjoint. For a subset $I \subseteq \mathbb{Z}$, we write

$$S(I) := \sum_{i \in I} \sigma_i, \quad M(I) := \max_{i \in I} \sigma_i.$$

Then define the events

$$\mathcal{A}_n := \left\{ M(I_n) > \varepsilon^{-3} a_n^{1/\alpha} \ell_\alpha(a_n), S(I_n \cup J_n) - M(I_n) < 2\varepsilon^{-1} a_n^{1/\alpha} \ell_\alpha(a_n) \right\}.$$

Importantly, we observe that the disjointness of the intervals involved in their definition readily yields that these events are independent. We use this fact in the proof of the following result.

Lemma 2.1. *As $n \rightarrow \infty$, it \mathbf{P} -a.s. holds that \mathcal{A}_n occurs infinitely often.*

Proof. By the independence of $(\mathcal{A}_n)_{n \geq 0}$ and the second Borel-Cantelli lemma, it will suffice to show that

$$(2.2) \quad \sum_{n=0}^{\infty} \mathbf{P}(\mathcal{A}_n) = \infty.$$

Since we have a continuous distribution for σ_0 , it holds that

$$\begin{aligned} \mathbf{P}(\mathcal{A}_n) &= \sum_{i \in I_n} \mathbf{P}(\mathcal{A}_n, x_n = i) \\ &= (a_n - b_{n-1}) \mathbf{P}\left(\sigma_{a_n} > \varepsilon^{-3} a_n^{1/\alpha} \ell_\alpha(a_n), S(I_n \cup J_n \setminus \{a_n\}) < 2\varepsilon^{-1} a_n^{1/\alpha} \ell_\alpha(a_n)\right) \\ &\geq (a_n - b_{n-1}) \mathbf{P}\left(\sigma_0 > \varepsilon^{-3} a_n^{1/\alpha} \ell_\alpha(a_n)\right) \mathbf{P}\left(S_{2b_n} < 2\varepsilon^{-1} a_n^{1/\alpha} \ell_\alpha(a_n)\right). \end{aligned}$$

From (1.2) it is easy to check that the first of these probabilities satisfies

$$(a_n - b_{n-1}) \mathbf{P}\left(\sigma_0 > \varepsilon^{-3} a_n^{1/\alpha} \ell_\alpha(a_n)\right) \sim \varepsilon^{3\alpha} \ell_\alpha(a_n)^{-\alpha}$$

as $n \rightarrow \infty$. To control the second probability, we treat the cases $\alpha \in (0, 1)$ and $\alpha = 1$ separately. In the case $\alpha \in (0, 1)$, it is well-known that (see [11, p. 62, Table 2.1], for example), as $n \rightarrow \infty$, $n^{-1/\alpha} S_n \Rightarrow \Xi_\alpha$ in distribution, where Ξ_α is a random variable with a maximally-asymmetric α -stable law supported on $(0, \infty)$. Hence

$$\mathbf{P} \left(S_{2b_n} < 2\varepsilon^{-1} a_n^{1/\alpha} \right) \rightarrow \mathbf{P} \left(\Xi_\alpha < (2\varepsilon^{-1})^{1-1/\alpha} \right)$$

as $n \rightarrow \infty$. Thus we find that

$$\mathbf{P}(\mathcal{A}_n) \rightarrow \varepsilon^{3\alpha} \mathbf{P} \left(\Xi_\alpha < (2\varepsilon^{-1})^{1-1/\alpha} \right) > 0,$$

and the result at (2.2) follows. In the case $\alpha = 1$, it is instead the case that (again, see [11, p. 62, Table 2.1], for example), as $n \rightarrow \infty$, $n^{-1}(S_n - n \log n) \Rightarrow \Xi_1$ in distribution, where Ξ_1 is a random variable with a maximally-symmetric 1-stable law supported on \mathbb{R} . Hence

$$\mathbf{P} \left(S_{2b_n} < 2\varepsilon^{-1} a_n \log a_n \right) \rightarrow \mathbf{P} \left(\Xi_1 < -\log(2/\varepsilon) \right) > 0$$

as $n \rightarrow \infty$. Thus we find that, as $n \rightarrow \infty$,

$$\mathbf{P}(\mathcal{A}_n) \sim c \log(a_n)^{-1} \sim c/(2n \log n),$$

for some constant $c > 0$, and so the result at (2.2) also follows in this case. \square

Lemma 2.2. *As $n \rightarrow \infty$, it \mathbf{P} -a.s. holds that \mathcal{E}_n occurs infinitely often.*

Proof. We start by defining the event

$$\mathcal{B}_n := \left\{ \bar{S}_{b_{n-1}} < a_n^{1/\alpha} \ell_\alpha(a_n) \right\},$$

which we claim holds eventually, \mathbf{P} -a.s. To prove this, note that it is an elementary computation to deduce from (1.2) that

$$1 - \mathbf{E} \left(e^{-\theta \sigma_0} \right) \sim c_\alpha \theta^\alpha \ell_\alpha(\theta^{-1})$$

as $\theta \rightarrow 0$, where

$$c_\alpha := \begin{cases} \Gamma(1 - \alpha), & \text{if } \alpha \in (0, 1), \\ 1, & \text{if } \alpha = 1, \end{cases}$$

with Γ the usual gamma function. Thus, for any c_n such that

$$\lim_{n \rightarrow \infty} n c_n^{-\alpha} \ell_\alpha(c_n) = 0,$$

applying Markov's inequality we have,

$$\begin{aligned} \mathbf{P}(S_n > c_n) &\leq \frac{1 - \mathbf{E}\left(e^{-c_n^{-1} S_n}\right)}{1 - e^{-1}} \\ &= (1 - e^{-1})^{-1} \left(1 - \mathbf{E}\left(e^{-c_n^{-1} \sigma_0}\right)^n\right) \\ &\sim c_\alpha (1 - e^{-1})^{-1} n c_n^{-\alpha} \ell_\alpha(c_n). \end{aligned}$$

Finally note that there is a constant $c > 0$ such that, as $n \rightarrow \infty$, eventually

$$\frac{(2b_{n-1} + 1) \ell_\alpha(a_n^{1/\alpha} \ell_\alpha(a_n))}{a_n \ell_\alpha(a_n)^\alpha} < c n^{-2} \rightarrow 0,$$

where we have used the fact that $a_{n-1}/a_n \sim n^{-2}$. Hence we deduce that there exists a constant $c > 0$ such that, as $n \rightarrow \infty$, eventually

$$\mathbf{P}\left(\bar{S}_{b_{n-1}} > a_n^{1/\alpha} \ell_\alpha(a_n)\right) = \mathbf{P}\left(S_{2b_{n-1}+1} > a_n^{1/\alpha} \ell_\alpha(a_n)\right) < c n^{-2},$$

and by the Borel-Cantelli lemma the claim is proved.

Now, it is a consequence of Lemma 2.1 and the conclusion of the previous paragraph that $\mathcal{A}_n \cap \mathcal{B}_n$ occurs infinitely often, \mathbf{P} -a.s. Thus to complete the proof it will suffice to show that $\mathcal{A}_n \cap \mathcal{B}_n \subseteq \mathcal{E}_n$. However, this is straightforward, since on $\mathcal{A}_n \cap \mathcal{B}_n$, we have that

$$M_{a_n} = \max\{M_{b_{n-1}}, M(I_n)\} \geq M(I_n) > \varepsilon^{-3} a_n^{1/\alpha} \ell_\alpha(a_n),$$

and also

$$\bar{S}_{b_n} - M_{a_n} \leq \bar{S}_{b_{n-1}} + S(I_n \cup J_n) - M(I_n) < (2\varepsilon^{-1} + 1) a_n^{1/\alpha} \ell_\alpha(a_n)$$

as required, since $\varepsilon < 1$. \square

We now proceed to establish that, on the event \mathcal{E}_n , at the time t_n the BTM is completely localised, up to an ε -dependent error, on the site x_n . We first state a general localisation result that is valid for arbitrary times $t > 0$ and deterministic trapping landscapes σ , before specialising to the time t_n and invoking the properties of the trapping landscape contained in \mathcal{E}_n .

Lemma 2.3. *Let σ be a deterministic strictly-positive trapping landscape, and denote by X the continuous-time \mathbb{Z} -valued Markov chain, started from 0, with transition rates given by (1.1), with P_σ its law. For a site $z \in \mathbb{Z}$ denote by τ_z the first hitting time by X of z , i.e. $\tau_z := \inf\{s > 0 : X_s = z\}$. Then, for each pair of sites $0 < x < y$ and each time $t > 0$,*

$$P_\sigma(\tau_x < t < \tau_y \wedge \tau_{-y}) \geq \left(\frac{y}{x+y} - t^{-1} x \sum_{-y < z < x} \sigma_z\right) \left(1 - \frac{t}{(y-x)\sigma_x}\right)$$

and

$$P_\sigma(X_s = x | \tau_x < t < \tau_y \wedge \tau_{-y}) \geq 1 - \left(1 - \frac{t}{(y-x)\sigma_x}\right)^{-1} \left(1 - \frac{\sigma_x}{\sum_{-y \leq z \leq y} \sigma_z}\right).$$

Proof. We prove Lemma 2.3 in a similar manner to the equivalent results in [5], which were established in the one-sided case. Here we adapt these arguments to the two-sided case.

Consider the first statement. Applying the Markov property at time τ_x , we have that

$$(2.3) \quad \begin{aligned} P_\sigma(\tau_x < t < \tau_y \wedge \tau_{-y}) &\geq P_\sigma(\tau_x \wedge \tau_{-y} < t, \tau_x < \tau_{-y}) P_\sigma(\tau_{-y} \wedge \tau_y - \tau_x > t) \\ &\geq (P_\sigma(\tau_x < \tau_{-y}) - P_\sigma(\tau_x \wedge \tau_{-y} \geq t)) P_\sigma(\tau_{-y} \wedge \tau_y - \tau_x > t). \end{aligned}$$

By basic properties of random walks, it holds that $P_\sigma(\tau_x < \tau_{-y}) = \frac{y}{x+y}$. The second and third probabilities in (2.3) are simple to bound using the methods of [5]. For the second term, we can apply the following upper bound (see [5, Proposition 2.1])

$$P_\sigma(\tau_x \wedge \tau_{-y} \geq t) \leq t^{-1} x \sum_{-y < z < x} \sigma_z.$$

Similarly, for the third term we can apply the following lower bound (see [5, Proposition 2.2])

$$(2.4) \quad P_\sigma(\tau_{-y} \wedge \tau_y - \tau_x \leq t) \leq \frac{t}{(y-x)\sigma_x}.$$

Combining the above bounds establishes the first statement.

Consider then the second statement. Define the continuous-time $[-y, y]$ -valued Markov chain \hat{X} , started from x , with transition rates given by (1.1) (interpreting \sim as denoting nearest neighbours on $[-y, y] \cap \mathbb{Z}$), and let \hat{P}_σ be its law. Then it is clear that $(\hat{X}_{s \wedge \tau_y \wedge \tau_{-y}})_{s \geq 0}$ has the same distribution as $(X_{(s+\tau_x) \wedge \tau_y \wedge \tau_{-y}})_{s \geq 0}$. Hence, applying the Markov property at τ_x , we have that

$$(2.5) \quad \begin{aligned} P_\sigma(X_t \neq x | \tau_x < t < \tau_y \wedge \tau_{-y}) &\leq \sup_{s \leq t} \hat{P}_\sigma(\hat{X}_s \neq x | \tau_y \wedge \tau_{-y} > t) \\ &\leq \frac{\sup_{s \leq t} \hat{P}_\sigma(\hat{X}_s \neq x)}{P_\sigma(\tau_y \wedge \tau_{-y} - \tau_x > t)}. \end{aligned}$$

The denominator of (2.5) has already been estimated at (2.4). For the numerator, note that the process \hat{X} satisfies the following localisation bound (see [5, Proposition 2.3])

$$\inf_{t \geq 0} \hat{P}_\sigma(\hat{X}_t = x) \geq \frac{\sigma_x}{\sum_{-y \leq z \leq y} \sigma_z},$$

and the result is proved. \square

Applying the previous result together with the definitions of \mathcal{E}_n , t_n and x_n , we readily deduce the following.

Corollary 2.4. *On the event \mathcal{E}_n , it holds that*

$$(2.6) \quad P_\sigma(X_{t_n} = x_n) > \left(\frac{1}{1+\varepsilon} - 3\varepsilon \right) \left(1 - \frac{\varepsilon^2}{1-\varepsilon} \right) \left(1 - \left(1 - \frac{\varepsilon^2}{1-\varepsilon} \right)^{-1} 3\varepsilon^2 \right).$$

The right-hand side of (2.6) can be written $1 - O(\varepsilon)$ as $\varepsilon \rightarrow 0$.

Observe that, for any $\varepsilon \in (0, 1)$, the times $t_n \rightarrow \infty$ as $n \rightarrow \infty$. Hence, since $\varepsilon \in (0, 1)$ was arbitrary, as a simple consequence of Lemma 2.2 and Corollary 2.4 we obtain the lim sup part of Theorem 1.1. Furthermore, note that on \mathcal{E}_n we have that

$$0 \leq \frac{S_{a_n}}{M_{a_n}} - 1 = \frac{S_{a_n} - M_{a_n}}{M_{a_n}} \leq \frac{\bar{S}_{a_n} - M_{a_n}}{M_{a_n}} < 3\varepsilon^2.$$

Hence, since $a_n \rightarrow \infty$ and $\varepsilon \in (0, 1)$ was arbitrary, we also deduce from Lemma 2.2 the lim inf part of Theorem 1.2.

§ 3. Complete delocalisation

In this section we prove that the BTM is completely delocalised at arbitrarily large times, that is, we prove the lim inf part of Theorem 1.1. As in Section 2, our approach is to show that certain favourable configurations of the trapping landscape occur infinitely often \mathbf{P} -a.s., and moreover, when such favourable configurations arise, the BTM is highly delocalised.

Throughout this section we again introduce an $\varepsilon \in (0, 1)$ to act as our error threshold. We also introduce a parameter $K \in \mathbb{N}$ which measures the extent of the spread of the probability mass function of the BTM. Again, unless explicitly stated, both ε and K will remain fixed throughout this section, and as such we suppress the explicit dependence of other notation on ε and K .

To define the favourable configurations, recall the scale a_n from (2.1) and further define, for $k \in \mathbb{Z}$, the evenly spaced sites $a_{n,k} := ka_n$, the corresponding intervals

$$I_{n,k} := [a_{n,k}, a_{n,k+1}),$$

and the events, for $n \in \mathbb{N}$,

$$\mathcal{E}_n := \bigcap_{k \in [-K, K]} \left\{ S(I_{n,k}) \in \left(\frac{1}{2} a_n^{1/\alpha} \ell_\alpha(a_n), 2 a_n^{1/\alpha} \ell_\alpha(a_n) \right), M(I_{n,k}) < \varepsilon a_n^{1/\alpha} \ell_\alpha(a_n) \right\}.$$

We will show that, for each $\varepsilon \in (0, 1)$ and $K \in \mathbb{N}$, the events \mathcal{E}_n occur infinitely often \mathbf{P} -a.s. (see Lemma 3.2). Moreover, we show that, when the event \mathcal{E}_n occurs, at the time

$$t_n := 12a_n^{1+1/\alpha} \ell_\alpha(a_n)$$

no site carries a non-negligible ((ε, K) -dependent) proportion of the probability mass of the BTM (see Lemma 3.4).

To establish that the events \mathcal{E}_n occur infinitely often, we again start by proving the corresponding result for a closely related sequence of independent events. In particular, define the collection of intervals $(\tilde{I}_{n,k})_{n \geq 0, k \in \mathbb{N}}$ by setting

$$\tilde{I}_{n,0} := [(K+1)a_{n-1}, a_n), \quad \tilde{I}_{n,-1} := [a_n, -Ka_{n-1})$$

and $\tilde{I}_{n,k} = I_{n,k}$ for $k \in [-K, K] \setminus \{-1, 0\}$. We then set

$$\mathcal{A}_n := \bigcap_{k \in [-K, K]} \left\{ S(\tilde{I}_{n,k}) \in \left(\frac{1}{2} a_n^{1/\alpha} \ell_\alpha(a_n), \frac{3}{2} a_n^{1/\alpha} \ell_\alpha(a_n) \right), M(\tilde{I}_{n,k}) < \varepsilon a_n^{1/\alpha} \ell_\alpha(a_n) \right\}.$$

Observe that the intervals $(\tilde{I}_{n,k})_{n \geq 0, k \in [-K, K]}$ are distinct for sufficiently large n . Hence, as $n \rightarrow \infty$ the events \mathcal{A}_n are eventually independent. We use this fact in the proof of the following result.

Lemma 3.1. *As $n \rightarrow \infty$, it \mathbf{P} -a.s. holds that \mathcal{A}_n occurs infinitely often.*

Proof. By the eventual independence of $(\mathcal{A}_n)_{n \geq 0}$ and the second Borel-Cantelli lemma, it suffices to show that $\sum_{n=0}^{\infty} \mathbf{P}(\mathcal{A}_n) = \infty$. It is well-known ([10]) that as $n \rightarrow \infty$

$$\left(n^{-1/\alpha} \ell_\alpha(n)^{-1} S_{\lfloor nt \rfloor} \right)_{t \geq 0} \Rightarrow (\Xi_\alpha(t))_{t \geq 0}$$

weakly in the space $D(\mathbb{R}^+)$ of real-valued càdlàg functions equipped with the standard Skorohod J_1 topology (see [12] for the definition), where $\Xi_\alpha(t)$ denotes an α -stable subordinator for $\alpha \in (0, 1)$, and $\Xi_1(t) := t$. Notice that the functionals $f \mapsto f(1)$ and $f \mapsto \sup_{t \in [0, 1]} \Delta f(t)$, where $\Delta f(t) := f(t) - f(t^-)$ denotes the jump in the function f at the time t , are both continuous in the J_1 topology for functions that are continuous at $t = 1$, and moreover that all fixed times are continuity times for Ξ_α . Hence it follows that

$$\mathbf{P}(\mathcal{A}_n) \rightarrow \mathbf{P} \left(\Xi_\alpha(1) \in (1/2, 3/2), \sup_{t \in [0, 1]} \Delta \Xi_\alpha(t) < \varepsilon \right)^{2K+1} > 0$$

as desired. \square

Lemma 3.2. *As $n \rightarrow \infty$, it \mathbf{P} -a.s. holds that \mathcal{E}_n occurs infinitely often.*

Proof. We start by defining the event

$$\mathcal{B}_n := \left\{ S(-Ka_{n-1}, (K+1)a_{n-1}) < \left(\frac{1}{2} \wedge \varepsilon\right) a_n^{1/\alpha} \ell_\alpha(a_n) \right\},$$

which we claim holds eventually, \mathbf{P} -a.s. To prove this, recall from the proof of Lemma 2.2 that, for any c_n such that $\lim_{n \rightarrow \infty} nc_n^{-\alpha} \ell_\alpha(c_n) = 0$, there exists a constant $c > 0$ such that, as $n \rightarrow \infty$, eventually $\mathbf{P}(S_n > c_n) < cnc_n^{-\alpha} \ell_\alpha(c_n^{-1})$. Now consider that there is a constant $c > 0$ such that, as $n \rightarrow \infty$, eventually

$$\frac{a_{n-1} \ell_\alpha\left(\left(\frac{1}{2} \wedge \varepsilon\right) a_n^{1/\alpha} \ell_\alpha(a_n)\right)}{a_n \ell_\alpha(a_n)^\alpha} < cn^{-2} \rightarrow 0,$$

where we use the fact that $a_{n-1}/a_n \sim n^{-2}$. Hence we deduce that there exists a constant $c > 0$ such that, as $n \rightarrow \infty$, eventually

$$\begin{aligned} & \mathbf{P} \left(S(-Ka_{n-1}, (K+1)a_{n-1}) < \left(\frac{1}{2} \wedge \varepsilon\right) a_n^{1/\alpha} \ell_\alpha(a_n) \right) \\ &= \mathbf{P} \left(S_{2(K+1)a_{n-1}} > \left(\frac{1}{2} \wedge \varepsilon\right) a_n^{1/\alpha} \ell_\alpha(a_n) \right) < cn^{-2}, \end{aligned}$$

and by the Borel-Cantelli lemma the claim is proved.

Now, it is a consequence of Lemma 3.1 and the previous paragraph that $\mathcal{A}_n \cap \mathcal{B}_n$ occurs infinitely often, \mathbf{P} -a.s. Thus to complete the proof it will suffice to show that $\mathcal{A}_n \cap \mathcal{B}_n \subseteq \mathcal{E}_n$. However, this is straightforward, since on $\mathcal{A}_n \cap \mathcal{B}_n$ each $k = -1, 0$ satisfies

$$M(I_{n,k}) \leq \max \left\{ M(\tilde{I}_{n,k}), S(-Ka_{n-1}, (K+1)a_{n-1}) \right\} < \varepsilon a_n^{1/\alpha} \ell_\alpha(a_n),$$

and

$$S(I_{n,k}) \in \left(S(\tilde{I}_{n,k}), S(\tilde{I}_{n,k}) + S(-Ka_{n-1}, (K+1)a_{n-1}) \right) \in (1/2, 2) a_n^{1/\alpha} \ell_\alpha(a_n),$$

and moreover the conditions on $M(I_{n,k})$ and $S(I_{n,k})$ for each $k \in [-K, K] \setminus \{-1, 0\}$ are identical in the events \mathcal{A}_n and \mathcal{E}_n . \square

We now proceed to study the behaviour of the BTM on the event \mathcal{E}_n . In particular, our first aim is to show that at time t_n no site in the interval $[a_{n,-K+1}, a_{n,K}]$ carries significant mass.

Lemma 3.3. *If \mathcal{E}_n holds, then*

$$\sup_{x \in [a_{n,-K+1}, a_{n,K}]} P_\sigma(X_{t_n} = x) < 4\varepsilon.$$

Proof. We first introduce the notation

$$V(x, r) := S([x - r + 1, x + r - 1])$$

and denote the quenched heat kernel of the Markov chain X by

$$(3.1) \quad p_t^\sigma(x, y) = \sigma_y^{-1} P_\sigma(X_t = y | X_0 = x).$$

Throughout the proof we suppose that \mathcal{E}_n holds, and we note that on this event

$$(3.2) \quad \frac{1}{2} a_n^{1/\alpha} \ell_\alpha(a_n) < V(x, a_n) < 6 a_n^{1/\alpha} \ell_\alpha(a_n), \quad \forall x \in [a_{n, -K+1}, a_{n, K}].$$

Next, in this one-dimensional setting, it is possible to check by applying the argument of [9, Proposition 4.1] that $p_{2a_n V(x, a_n)}^\sigma(x, x) \leq 2/V(x, a_n)$. Hence, since $(p_t^\sigma(x, x))_{t \geq 0}$ is decreasing in t , it follows from (3.2) that

$$(3.3) \quad p_{t_n}^\sigma(x, x) < \frac{4}{a_n^{1/\alpha} \ell_\alpha(a_n)}, \quad \forall x \in [a_{n, -K+1}, a_{n, K}].$$

In conjunction with the Cauchy-Schwarz bound

$$p_t^\sigma(0, x) \leq \sqrt{p_t^\sigma(0, 0) p_t^\sigma(x, x)}, \quad \forall x \in \mathbb{Z}, t > 0,$$

and the estimate $\max_{x \in [a_{n, -K+1}, a_{n, K}]} \sigma_x < \varepsilon a_n^{1/\alpha}$ that holds on \mathcal{E}_n , (3.3) implies the result. \square

To extend the bound of the previous lemma to hold uniformly over the entire integer lattice, we will check that on \mathcal{E}_n the Markov chain X does not exit the interval $[a_{n, -K+1}, a_{n, K}]$ quickly. To this end, we introduce the notation

$$\tau(x, r) := \inf\{t \geq 0 : |X_t - x| \geq r\}.$$

Lemma 3.4. *If \mathcal{E}_n holds, then*

$$\sup_{x \in \mathbb{Z}} P_\sigma(X_{t_n} = x) < 4\varepsilon + b_K,$$

where $(b_k)_{k \geq 1}$ is a deterministic sequence of positive numbers such that $b_k \rightarrow 0$ as $k \rightarrow \infty$.

Proof. In light of Lemma 3.3, to complete the proof it will suffice to show that on \mathcal{E}_n it holds that

$$P_\sigma(\tau(0, (K-1)a_n) \leq t_n) \leq b_K,$$

where $b_k := \mathbb{P}(\text{Bin}(\lfloor (k-1)/2 \rfloor, \frac{1}{32}) \leq 24)$, and $\text{Bin}(N, p)$ is a binomial random variable with parameters N and p .

Again, throughout the proof we suppose that \mathcal{E}_n holds. It is an elementary computation to check that $E_\sigma(\tau(x, r) | X_0 = y)$ is equal to

$$2 \sum_{i=x-r+1}^{x+r-1} \min \left\{ \frac{x+r-y}{x+r-i}, \frac{y-x+r}{i-x+r} \right\} \times \left(\frac{1}{i-x+r} + \frac{1}{x+r-i} \right)^{-1} \sigma_i$$

for every $x \in \mathbb{Z}$, $r \in \mathbb{N}$, $y \in (x-r, x+r)$. In particular, by the definition of \mathcal{E}_n we readily deduce from this that

$$E_\sigma(\tau(x, 2a_n) | X_0 = y) < 16a_n^{1+1/\alpha} \ell_\alpha(a_n)$$

and

$$E_\sigma(\tau(x, 2a_n) | X_0 = x) > a_n^{1+1/\alpha} \ell_\alpha(a_n)$$

for every $x \in [a_{n,-K+2}, a_{n,K-1}] \cap a_n \mathbb{Z}$, $y \in (x-2a_n, x+2a_n)$. Thus, applying the Markov property at time t , we deduce

$$\begin{aligned} a_n^{1+1/\alpha} \ell_\alpha(a_n) &< E_\sigma(\tau(x, 2a_n) | X_0 = x) \\ &\leq t + P_\sigma(\tau(x, 2a_n) \geq t | X_0 = x) 16a_n^{1+1/\alpha} \ell_\alpha(a_n) \end{aligned}$$

for every $x \in [a_{n,-K+2}, a_{n,K-1}] \cap a_n \mathbb{Z}$. In particular, setting $t = \tilde{t}_n := \frac{1}{2} a_n^{1+1/\alpha} \ell_\alpha(a_n)$, this implies

$$P_\sigma(\tau(x, 2a_n) \geq \tilde{t}_n | X_0 = x) > \frac{1}{32}$$

for every $x \in [a_{n,-K+2}, a_{n,K-1}] \cap a_n \mathbb{Z}$. Writing $\tau_0 = 0$ and

$$\tau_{i+1} := \inf \{t \geq \tau_i : X_t \in 2a_n \mathbb{Z} \setminus \{X_{\tau_i}\}\},$$

it follows that

$$\begin{aligned} P_\sigma(\tau(0, (K-1)a_n) \leq t_n) &\leq P_\sigma \left(\sum_{i=0}^{\lfloor (K-1)/2 \rfloor - 1} \tau_i \leq t_n \right) \\ &\leq P_\sigma \left(\sum_{i=0}^{\lfloor (K-1)/2 \rfloor - 1} \mathbf{1}_{\{\tau_i \geq \tilde{t}_n\}} \leq 24 \right) \\ &\leq \mathbb{P} \left(\text{Bin} \left(\lfloor (K-1)/2 \rfloor, \frac{1}{32} \right) \leq 24 \right) \end{aligned}$$

as desired. \square

Putting together Lemmas 3.2 and 3.4, recalling that $\varepsilon \in (0, 1)$ and $K \in \mathbb{N}$ were arbitrary, and noting that $t_n \rightarrow \infty$, we have thus established the lim inf part of Theorem 1.1. Furthermore, note that on \mathcal{E}_n we have that

$$\frac{S_{a_n}}{M_{a_n}} > \frac{1}{2\varepsilon}.$$

Hence we also deduce from Lemma 3.2 the lim sup part of Theorem 1.2.

§ 4. Delocalisation for traps with finite expectation

In this section, we deal with the case when $\mathbf{E}\sigma_0 < \infty$, and in particular establish (1.4). In this case, it is known that, for \mathbf{P} -a.e. realisation of the trapping environment, we have that

$$(4.1) \quad (\varepsilon X_{t/\varepsilon^2})_{t \geq 0} \rightarrow (B_t)_{t \geq 0}$$

in distribution in $D([0, \infty), \mathbb{R})$, where up to a deterministic linear time change $B = (B_t)_{t \geq 0}$ is a standard one-dimensional Brownian motion [1, Theorem 2.10]. Moreover, by the strong law of large numbers we \mathbf{P} -a.s. have that $\varepsilon \sum_{x \in \mathbb{Z}} \sigma_x \delta_{\varepsilon x} \rightarrow \mathbf{E}(\sigma_0)\lambda$ vaguely, where δ_x is the probability measure placing all its mass at x , and λ is Lebesgue measure on \mathbb{R} . It follows that we can apply the local limit theorem of [4, Theorem 1] to deduce that, \mathbf{P} -a.s.,

$$\lim_{t \rightarrow \infty} \max_{x \in \mathbb{Z}: |x| \leq R\sqrt{t}} \sqrt{t} \left| p_t^\sigma(0, x) - q(x/\sqrt{t}) \right| = 0,$$

where p^σ is the discrete heat kernel defined as at (3.1), and q is the density of B_1 with respect to $\mathbf{E}(\sigma_0)\lambda$. (For this, it is useful to note that in this setting the Euclidean metric coincides with the resistance metric, where we consider \mathbb{Z} as an electrical network with unit conductances between nearest neighbours. This means we can immediately apply [4, Proposition 14] to check the equicontinuity of the discrete heat kernels under the relevant scaling.) In particular, we obtain that, \mathbf{P} -a.s., for large t ,

$$\max_{x \in \mathbb{Z}: |x| \leq R\sqrt{t}} P_\sigma(X_t = x) \leq 2 \sup_{x \in \mathbb{R}: |x| \leq R} q(x) \max_{x \in \mathbb{Z}: |x| \leq R\sqrt{t}} \frac{\sigma_x}{\sqrt{t}}$$

Applying the strong law of large numbers again, one can readily check that the upper bound here converges to 0 for any $R < \infty$, \mathbf{P} -a.s. From (4.1), we also deduce that, \mathbf{P} -a.s.,

$$\sup_{x \in \mathbb{Z}: |x| > R\sqrt{t}} P_\sigma(X_t = x) \leq P_\sigma\left(\tau(0, R\sqrt{t}) < t\right) \rightarrow \mathbb{P}\left(\tau^B(0, R) < 1\right),$$

where $\tau^B(0, R)$ is the exit time of B from $(-R, R)$. Hence we find that

$$\limsup_{t \rightarrow \infty} \sup_{x \in \mathbb{Z}} P_\sigma(X_t = x) \leq \mathbb{P}\left(\tau^B(0, R) < 1\right)$$

for any $R < \infty$, \mathbf{P} -a.s. Since this bound can be made arbitrarily small by taking R large, we are done.

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